

Recall Cameron-Martin-Girsanov-Maruyama's formula:

- ▶ $(\Omega, \mathcal{F}, \mathcal{P}), (\mathcal{F}_t)_{t \geq 0}$: Probability space and its filtration
- ▶ $B = (B_t(\omega))_{t \in [0,1]}$: 1-dimensional (\mathcal{F}_t) -Brownian motion such that $B_0 = 0$.
- ▶ $\eta_t = \eta(t, \omega), t \in [0, 1]$: (\mathcal{F}_t) -adapted stochastic process and absolutely continuous in t s.t. $\eta_0 = 0$ and

$$\sup_{\omega \in \Omega} \int_0^1 \dot{\eta}_t(\omega)^2 dt < \infty.$$

[Theorem 18.1] One can define a probability measure Q on (Ω, \mathcal{F}) by

$$\frac{dQ}{dP}(\omega) = \exp \left\{ \int_0^1 \dot{\eta}_t dB_t - \frac{1}{2} \int_0^1 \dot{\eta}_t^2 dt \right\},$$

and under Q ,

$$\hat{B}_t := B_t - \eta_t, \quad t \in [0, 1]$$

is a Brownian motion.



To give a proof of Theorem 18.1, we prepare a lemma.

[Lemma 18.3] Assume that an (\mathcal{F}_t) -adapted continuous process $X = (X_t)_{t \geq 0}$ satisfies $X_0 = 0$ (a.s.) and

$$\varphi(X_t) - \frac{1}{2} \int_0^t \varphi''(X_s) ds : (\mathcal{F}_t)\text{-martingale}$$

for $\forall \varphi \in C_b^2(\mathbb{R})$. Then, X is an (\mathcal{F}_t) -Brownian motion. \square

☺ Take $\varphi(x) = e^{i\xi x}$, $\xi \in \mathbb{R}$ (or its real part and imaginary part separately). Then, by the condition, $e^{i\xi X_t} + \frac{\xi^2}{2} \int_0^t e^{i\xi X_s} ds$ is an (\mathcal{F}_t) -martingale. Therefore, for any \mathcal{F}_s -measurable and bounded function Φ on Ω , we have

$$\frac{d}{dt} E [e^{i\xi X_t} \Phi] = -\frac{\xi^2}{2} E [e^{i\xi X_t} \Phi], \quad t \geq s > 0$$

One can easily solve this ODE and obtain

$$E [e^{i\xi X_t} \Phi] = E [e^{i\xi X_s} \Phi] e^{-\frac{\xi^2}{2}(t-s)}$$

Take $e^{-i\xi X_s} \Phi$ for Φ and we obtain $E[e^{i\xi(X_t - X_s)} \Phi] = e^{-\frac{\xi^2}{2}(t-s)} E[\Phi]$.

This implies that $X_t - X_s \perp\!\!\!\perp \mathcal{F}_s$, $X_t - X_s \stackrel{\text{law}}{=} N(0, t - s)$.

Thus, the conclusion is shown. \square

[Remark] (1) If X is a Brownian motion, the property stated in Lemma 18.3 holds. Recall Proposition 16.1 with $\mathcal{L}_s = \frac{1}{2} \frac{\partial^2}{\partial x^2}$ for the Brownian motion.

(2) Lemma 18.3 claims that the **converse is also true**, or one can say that the **solution of the martingale problem for $\mathcal{L} = \frac{1}{2} \frac{\partial^2}{\partial x^2}$** is **unique** and it is the Wiener measure.

(3) By the characterization of the Brownian motion by martingale (**Lévy's theorem**), for X to be a Brownian motion, it is actually enough to check the above martingale property only for $\varphi(x) = x, x^2$. In other words, if X_t is a square integrable martingale and $X_t^2 - t$ is also a martingale (note that $\frac{1}{2}(x^2)'' = 1$), then X_t is an (\mathcal{F}_t) -Brownian motion. \square

[Proof of Theorem 18.1] [Step 1] Set

$$M_t := \exp \left\{ \int_0^t \dot{\eta}_s dB_s - \frac{1}{2} \int_0^t \dot{\eta}_s^2 ds \right\}, \quad t \in [0, 1]$$

under the probability measure P (i.e. we define $\int_0^t \dot{\eta}_s dB_s$ as a stochastic integral under P). Then, M_t is an (\mathcal{F}_t) -martingale.

☺ By Itô's formula,

$$\begin{aligned} dM_t &= M_t \left\{ \dot{\eta}_t dB_t - \frac{1}{2} \dot{\eta}_t^2 dt \right\} + \frac{1}{2} M_t \dot{\eta}_t^2 dt \\ &= M_t \dot{\eta}_t dB_t. \end{aligned}$$

In particular, M_t is written as a stochastic integral. Therefore, it is a martingale. □

The above proof is **a little rough**, since $f(x) = e^x \notin C_b^2$ so that we should be careful to apply Itô's formula. \rightarrow We need a **cutoff argument**. (Define $\varphi_n \in C_b^2(\mathbb{R})$ s.t. $\varphi_n(x) = e^x$, $|x| \leq n$ and $\sigma_n := \inf\{t > 0; |M_t| > n\}$. Show $\sigma_n \nearrow \infty$ (a.s.) and $M_{t \wedge \sigma_n}$ is uniformly integrable martingale by the assumption $\sup_{\omega \in \Omega} \int_0^1 \dot{\eta}_t(\omega)^2 dt < \infty$.) This shows M_t is **square integrable**.

Since M_t is a martingale, the measure Q on (Ω, \mathcal{F}) defined by $\frac{dQ}{dP} = M_1$ (i.e. $dQ = M_1 dP$) is a probability measure on Ω .

☺ The σ -additivity of Q is clear. It is a probability measure, since we have

$$Q(\Omega) = \int_{\Omega} M_1 dP = E^P[M_1] \underset{M:\text{martingale}}{=} E^P[M_0] \underset{M_0=1}{=} 1. \quad \square$$

[Step 2] Next, we show that $\hat{B}_t := B_t - \eta_t$ is a Brownian motion under Q . To this end, by Lemma 18.3, it is enough to show that

$$\varphi(\hat{B}_t) - \frac{1}{2} \int_0^t \varphi''(\hat{B}_s) ds$$

is Q -martingale for $\forall \varphi \in C_b^2(\mathbb{R})$.

Now, under P , noting that $dM_t = M_t \dot{\eta}_t dB_t$, by Itô's formula, we have

$$\begin{aligned}
 d\{M_t \cdot \varphi(\hat{B}_t)\} &= M_t d\{\varphi(\hat{B}_t)\} + \varphi(\hat{B}_t) dM_t + dM_t d\{\varphi(\hat{B}_t)\} \\
 &= M_t \varphi'(\hat{B}_t) \{dB_t - \dot{\eta}_t dt\} + \frac{1}{2} M_t \varphi''(\hat{B}_t) dt \\
 &\quad + \varphi(\hat{B}_t) M_t \dot{\eta}_t dB_t + M_t \dot{\eta}_t \varphi'(\hat{B}_t) dt \\
 &= dN_t + \frac{1}{2} M_t \varphi''(\hat{B}_t) dt,
 \end{aligned}$$

where

$$N_t := \varphi(0) + \int_0^t M_s \left\{ \varphi'(\hat{B}_s) + \dot{\eta}_s \varphi(\hat{B}_s) \right\} dB_s$$

is an (\mathcal{F}_t) -martingale under P , since it is a stochastic integral by a Brownian motion (under P). (Note that, since M_t is square integrable, the integrand $\in \mathcal{L}_T^2$.)

Writing this in an integrated form, we have

$$M_t \varphi(\hat{B}_t) = N_t + \frac{1}{2} \int_0^t M_s \varphi''(\hat{B}_s) ds.$$

Note that $M_0 \varphi(\hat{B}_0) = \varphi(0) = N_0$ holds at $t = 0$. In particular, for $t \geq s \geq 0$, we have

$$M_t \varphi(\hat{B}_t) - M_s \varphi(\hat{B}_s) = N_t - N_s + \frac{1}{2} \int_s^t M_r \varphi''(\hat{B}_r) dr. \quad (\star)$$

Using this, we will show for $1 \geq \forall t \geq \forall s \geq 0$, $\forall \mathcal{F}_s$ -measurable bounded function Φ :

$$E^Q[\varphi(\hat{B}_t)\Phi] = E^Q[\varphi(\hat{B}_s)\Phi] + E^Q \left[\frac{1}{2} \int_s^t \varphi''(\hat{B}_r) dr \cdot \Phi \right]. \quad (\star\star)$$

This implies that

$$\varphi(\hat{B}_t) - \frac{1}{2} \int_0^t \varphi''(\hat{B}_s) ds \quad \text{is a } (Q, (\mathcal{F}_t))\text{-martingale}$$

and, by Lemma 18.3, this concludes the proof of Theorem 18.1.

Finally, we show ($\star\star$):

$$\begin{aligned} E^Q[\varphi(\hat{B}_t)\Phi] &\stackrel{(1)}{=} E^P[M_1\varphi(\hat{B}_t)\Phi] \stackrel{(2)}{=} E^P[M_t\varphi(\hat{B}_t)\Phi] \\ &\stackrel{(\star)}{=} E^P \left[\left\{ M_s\varphi(\hat{B}_s) + N_t - N_s + \frac{1}{2} \int_s^t M_r\varphi''(\hat{B}_r) dr \right\} \Phi \right] \\ &\stackrel{(3)}{=} E^P[M_s\varphi(\hat{B}_s)\Phi] + \frac{1}{2} \int_s^t E^P[M_r\varphi''(\hat{B}_r)\Phi] dr \\ &\stackrel{(4)}{=} E^Q[\varphi(\hat{B}_s)\Phi] + E^Q \left[\frac{1}{2} \int_s^t \varphi''(\hat{B}_r) dr \cdot \Phi \right] \end{aligned}$$

Here,

- (1) follows by the definition of Q .
- (2): Noting that M_t is $(P, (\mathcal{F}_t))$ -martingale, we rewrite $E^P[M_1\varphi(\hat{B}_t)\Phi] = E^P[E^P[M_1|\mathcal{F}_t]\varphi(\hat{B}_t)\Phi] = E^P[M_t\varphi(\hat{B}_t)\Phi]$.
- (3) follows since N_t is a P -martingale.
- (4): Similarly to (2)+(1), we use the P -martingale property of M . □

[Application] (Construction of a weak solution of SDE)

Let $b(t, x)$ be a bounded Borel measurable function on $[0, 1] \times \mathbb{R}$ and consider 1-dimensional SDE:

$$dX_t = dB_t + b(t, X_t) dt, \quad (1)$$

where B_t is a Brownian motion under P . Then, set under P

$$M_t := \exp \left\{ \int_0^t b(s, B_s) dB_s - \frac{1}{2} \int_0^t b(s, B_s)^2 ds \right\}$$

$$dQ := M_1 dP.$$

Then, by taking $\eta_t = \int_0^t b(s, B_s) ds$ in Theorem 18.1, we see that

$$\hat{B}_t := B_t - \int_0^t b(s, B_s) ds \quad (2)$$

is a Brownian motion under Q .

However, (2) is rewritten as

$$dB_t = d\hat{B}_t + b(t, B_t) dt, \quad B_0 = 0.$$

In other words, B_t is a weak solution of the SDE (1) on a probability space (Ω, \mathcal{F}, Q) equipped with the Brownian motion \hat{B} . The method used here is called the **transform of drift**.

This method is useful also to show the **uniqueness of the weak solution**. i.e. for $i = 1, 2$, if $(\Omega^{(i)}, \mathcal{F}^{(i)}, P^{(i)})$, $(X^{(i)}, B^{(i)})$ are both solutions of the SDE (1), then the distributions of $X^{(1)}$ and $X^{(2)}$ on a path space $W^1 \equiv C([0, \infty), \mathbb{R})$ coincide. (one can take $[0, T]$ as the time interval instead of $[0, 1]$ for $\forall T > 0$.) (\rightarrow by similar transformation with $-b$ instead of b , one can reduce to Brownian motion. The details are omitted.)

§19 Ornstein-Uhlenbeck process — Example of SDE

- ▶ Consider SDE for $X_t \in \mathbb{R}^d$ with constant diffusion coefficient and linear drift term:

$$dX_t = \alpha dB_t + AX_t dt, \quad (3)$$

where α and A are $d \times d$ real matrices, and $B = (B_t)$ is a d -dimensional Brownian motion.

- ▶ In physics, (3) is a kind of (simplest) [Langevin equation](#).
- ▶ The SDE (3) can be solved in terms of the exponential function e^{tA} of matrices as follows:

$$X_t = e^{tA} X_0 + \int_0^t e^{(t-s)A} \alpha dB_s. \quad (4)$$

This is called [Duhamel's formula](#).

[Proof] One can check that X_t given by (4) is a solution of (3) by applying Itô's formula and noting $e^{0A} = I$ as follows:

$$\begin{aligned} dX_t & \stackrel{\text{Itô}}{=} Ae^{tA}X_0dt + e^{0A}\alpha dB_t + \left(\int_0^t Ae^{(t-s)A}\alpha dB_s \right) dt \\ & = AX_tdt + \alpha dB_t. \end{aligned}$$

The uniqueness of (strong) solution is easy, since the coefficients (constant and linear function) are Lipschitz continuous. □

- In particular, assuming that the initial value $X_0 = x \in \mathbb{R}^d$ is given, $X = (X_t)$ is a Gaussian process and for each $t \geq 0$,

$$X_t \stackrel{\text{law}}{=} N(m_t, V_t) \quad (5)$$

with $m_t = e^{tA}x \in \mathbb{R}^d$ and $V_t = \int_0^t e^{sA}\alpha\alpha^*e^{sA^*} ds \in \mathbb{R}^{d \times d}$, where α^*, A^* denote the transposed matrices of α, A .

[Proof] Since $E[(\text{stochastic integral})] = 0$ in (4), we easily see $E[X_t] = m_t$. To compute the covariance matrix V_t , take $\xi \in \mathbb{R}^d$ and see that

$$\begin{aligned} (V_t\xi, \xi) &= E\left[\left(\int_0^t e^{(t-s)A}\alpha dB_s, \xi\right)^2\right] \\ &= E\left[\left\{\int_0^t (dB_s, \alpha^* e^{(t-s)A^*} \xi)\right\}^2\right] \\ &\stackrel{\text{It\^o isometry}}{=} \int_0^t \left|\alpha^* e^{(t-s)A^*} \xi\right|^2 ds \quad (|\cdot| = \text{Euclidean norm}) \\ &= \int_0^t (\alpha^* e^{sA^*} \xi, \alpha^* e^{sA^*} \xi) ds = \int_0^t (e^{sA}\alpha\alpha^* e^{sA^*} \xi, \xi) ds. \quad \square \end{aligned}$$

- In particular from (5), if V_t is positive definite (e.g. if $\alpha = 0$, then $V_t = 0$ so that degenerate), we have

$$P_x(X_t \in dy) = \frac{1}{(2\pi)^{d/2}(\det V_t)^{1/2}} e^{-\frac{1}{2}(x-m_t, V_t^{-1}(x-m_t))} dy.$$

This is called **Mehler's formula**.

- Let us study the **limit distribution of X_t as $t \rightarrow \infty$** . For simplicity, assume two conditions:

- $[\alpha, A] = 0$, i.e. $\alpha A - A\alpha = 0$, α and A are commutative.
- A is a symmetric matrix and $A < 0$ (i.e. negative definite).

Then, the distribution of X_t converges weakly as $t \rightarrow \infty$ to $N(0, \alpha\alpha^*(-2A)^{-1})$ which is independent of the starting point x of X_t .

☺ By $A < 0$, we have $m_t \rightarrow 0$. For the covariance, noting $[\alpha, A] = [\alpha^*, A^*] = [\alpha^*, A] = 0$, we have

$$\begin{aligned} V_t &= \int_0^t \alpha \alpha^* e^{2sA} ds \\ &= \alpha \alpha^* (2A)^{-1} (e^{2tA} - I) \rightarrow \alpha \alpha^* (-2A)^{-1}. \quad \square \end{aligned}$$

- In fact, in general, instead of the above two conditions, if real parts of all eigenvalues of A are negative, noting $\|e^{sA}\| \leq Ce^{-cs}$, $c, C > 0$, one can show that

$$V_t \rightarrow V_\infty := \int_0^\infty e^{sA} \alpha \alpha^* e^{sA^*} ds,$$

see Karatzas-Shreve p.357, (6.19), or Lunardi Trans. AMS, 1997.

[Definition 19.1] We call the solution $X = (X_t)_{t \geq 0}$ of SDE (or Markov process in general) **ergodic** or **positive recurrent**, if the distribution of X_t converges weakly as $t \rightarrow \infty$ to a certain probability distribution independently of the initial distribution (distribution of X_0). □

[Remark] Original definition of ergodicity is that the sample mean (time average) converges to the ensemble mean (average under a certain probability measure μ), that is, for $\forall \varphi \in C_b(\mathbb{R}^d)$,

$$\frac{1}{T} \int_0^T \varphi(X_t) dt \xrightarrow{T \rightarrow \infty} \int_{\mathbb{R}^d} \varphi(x) \mu(dx), \quad a.s.\omega.$$

Probabilistically, this corresponds to the strong law of large numbers (LLN). □

[Remark] The distribution of the Brownian motion

$$\frac{1}{(2\pi t)^{d/2}} e^{-\frac{|x|^2}{2t}} dx$$

spreads and, differently from the Ornstein-Uhlenbeck process, does not converge to a probability measure as $t \rightarrow \infty$.

Brownian motion is

- null recurrent when $d = 1, 2$
- transient (not recurrent) when $d \geq 3$.

“recurrent” means that X_t hits any open set (a.s.).

“null” means that the limit distribution is zero measure.